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Geometry of the Entropic Trust Region for Maximally Dense Crystallographic Symmetry Group Packings



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Stochastic relaxation is a well-known approach used to solve problems in machine learning and artificial intelligence, particularly when dealing with complicated optimization landscapes. Here, our focus is on identifying maximally dense packings of compact sets into n-dimensional Euclidean space, specifically packings restricted to the Crystallographic Symmetry Groups (CSGs). Utilizing stochastic relaxation, we have developed a non-Euclidean trust region algorithm called the Entropic Trust Region Packing Algorithm (ETRPA). The ETRPA is a variant of the natural gradient learning approach, equipped with adaptive selection quantile fitness rewriting. Since CSGs induce a toroidal topology on the configuration space, the ETRPA's search is performed on a statistical manifold of Extended Multivariate von Mises (EMvM) probability distributions, a parametric family of probability measures defined on an n-dimensional torus. To gain insight into the geometric properties of ETRPA, we establish a connection with the generalized proximal method. This connection allows us to examine the algorithm's behavior using local dual geodesic flows that maximize the stochastic dependence among elements of the EMvM distributed random vector. Consequently, ETRPA's theoretical foundation in evolutionary dynamics, statistical physics, and recurrent neural computing can be interpreted in terms of group equivariant geometric learning, providing a deeper understanding of the algorithm and its application in material science.

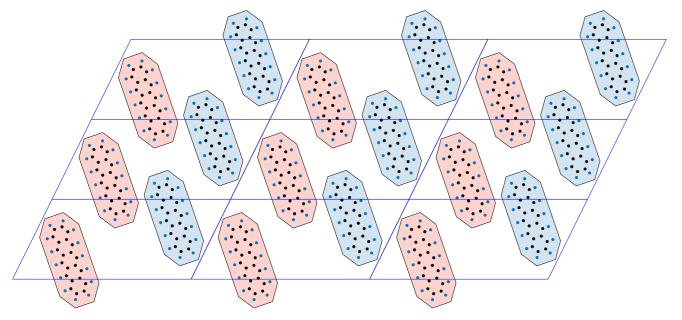
Problem statement

Given K a compact subset of \mathbb{R}^n and G a Crystallographic Symmetry Group (CSG) - a discrete group of isometries of \mathbb{R}^n containing a lattice subgroup - determine the configuration \mathcal{K}_G consisting only of non-overlapping copies of K generated as an orbit under the action of G on \mathbb{R}^n , such that the ratio of the filled to the whole space is maximized over the whole equivalence class \mathcal{G} of G. Formally,

$$\mathcal{K}_{\max} = \underset{\mathcal{K}_G: G \in \mathcal{G}}{\operatorname{argmax}} \rho\left(\mathcal{K}_G\right), \ \mathcal{G} = \{H | H \cong G\}.$$

$$(1)$$

where $0 \le \rho(\cdot) \le 1$ is the packing density. Our work is motivated by the problem of Crystal Structure Prediction, in which given some molecular shape, the goal is to predict a synthesizable periodic structure.



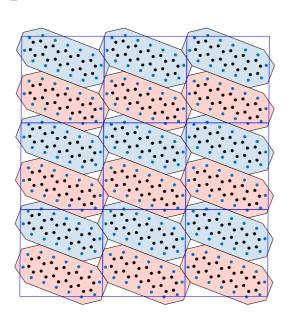


Figure 1: 2-dimensional CSG packings, wherein \mathcal{G} is the p2 plane group, and K is a geometric representation of pentacene by the convex hull of the atomic positions of the molecule, with an offset given by hydrogen's van der Waals radius of 1.09Å. (Left) Packing configuration with a density of 0.4990761. (Right) Packing configuration with a density of 0.9533821. The blue lines denote the crystal lattice, while the colors denote symmetry operations represented by cosets gL, where $g \in G$ is an element of the p2 plane, and L is the lattice subgroup of G. The dots symbolize the atomic positions of hydrogen (blue) and carbon (black).

Etropic Trust Region Packing Algorithm (ETRPA)

Stochastic relaxation and the entropic trust region

We transform Eq. (1) via stochastic relaxation [3] to the problem of $\tilde{\boldsymbol{\theta}} = \operatorname{argmax}_{\boldsymbol{\theta} \in \boldsymbol{\Theta}} J(\boldsymbol{\theta})$, with $J(\boldsymbol{\theta}) := E[\mathbf{F}|\boldsymbol{\theta}] = \int_{\mathcal{X}} \mathbf{F}(\mathbf{x}) dP(\boldsymbol{\theta})$ being the expected fitness \mathbf{F} of the packing density ρ under some probability measure $dP(\boldsymbol{\theta})$ from family of probability measures $S = \{dP(\boldsymbol{\theta}) \mid \boldsymbol{\theta} \in \boldsymbol{\Theta} \subseteq \mathbb{R}^n\}$ defined on some configuration space \mathcal{X} and solve it using first order trust region method where the trust region radius is given by the Kullback-Leibler divergence from $P_{\boldsymbol{\theta}}$ to $P_{\boldsymbol{\theta}+\delta\boldsymbol{\theta}}$, defined by $D_{KL}(P_{\boldsymbol{\theta}} \mid\mid P_{\boldsymbol{\theta}+\delta\boldsymbol{\theta}}) = \int_{\mathcal{X}} \ln\left(\frac{dP(\boldsymbol{\theta})}{dP(\boldsymbol{\theta}+\delta\boldsymbol{\theta})}\right) dP(\boldsymbol{\theta})$. The update equation reads

$$\boldsymbol{\theta}^{t+1} = \boldsymbol{\theta}^t + \Delta^t \frac{\widetilde{\nabla} J(\boldsymbol{\theta}^t)}{\parallel \widetilde{\nabla} J(\boldsymbol{\theta}^t) \parallel_{\mathcal{I}_{\boldsymbol{\theta}^t}}}, \tag{2}$$

where $\widetilde{\nabla} J(\boldsymbol{\theta}) = \mathcal{I}_{\boldsymbol{\theta}}^{-1} \nabla_{\boldsymbol{\theta}} J(\boldsymbol{\theta})$ is the natural gradient [1], $\|\cdot\|_{\mathcal{I}_{\boldsymbol{\theta}}}$ is the norm associated with the inner product induced by the Fisher metric tensor $\mathcal{I}_{\boldsymbol{\theta}}$ with elements $\mathcal{I}_{\boldsymbol{\theta} ij} = \int_{\mathcal{X}} \frac{\partial \ln(p(\boldsymbol{\theta}))}{\partial \theta_i} \frac{\partial \ln(p(\boldsymbol{\theta}))}{\partial \theta_j} dP(\boldsymbol{\theta})$, $p(\boldsymbol{\theta}) = \frac{dP(\boldsymbol{\theta})}{d\boldsymbol{\nu}}$ is the Radon-Nikodym derivative with respect to a reference measure $\boldsymbol{\nu}$ on \mathcal{X} and $\nabla_{\boldsymbol{\theta}}$ is the Euclidean gradient in $\boldsymbol{\theta}$ -coordinates.

Adaptive selection quantile

We implement a truncation selection transformation for $q \geq 1$ of the fitness function ${\bf F}$

$$q \mathbf{1}_{\mathbf{F}_{1-\frac{1}{q}}^{\tilde{\boldsymbol{\theta}}}}(\mathbf{x}) = \begin{cases} q & \text{if } \mathbf{F}(\mathbf{x}) \ge \mathbf{F}_{1-\frac{1}{q}}^{\tilde{\boldsymbol{\theta}}}, \\ 0 & \text{otherwise,} \end{cases}$$

where $\mathbf{F}_{1-\frac{1}{q}}^{\boldsymbol{\theta}}$ is the $P_{\tilde{\boldsymbol{\theta}}}$'s (q-1)th q-quantile of the fitness \mathbf{F} and $\mathbf{1}.(\cdot)$ is the indicator function. Assuming $P_{\tilde{\boldsymbol{\theta}}}$ is from the exponential family statistical model then, for a fixed q^*

$$\nabla_{\boldsymbol{\theta}} J(\boldsymbol{\theta}) \mid_{\boldsymbol{\theta} = \tilde{\boldsymbol{\theta}}} = \boldsymbol{\mu}_{\mathbf{F}_{1-\frac{1}{q^*}}} - \boldsymbol{\mu},$$
 (3)

where $\boldsymbol{\mu}_{\mathbf{F}_{1-\frac{1}{q^*}}} = \int_{\mathbf{F}_{1-\frac{1}{*}}}^{\infty} \mathbf{F}^{-1}(y) \exp \left\{ \tilde{\boldsymbol{\theta}}^{\mathsf{T}} \mathbf{t} \left(\mathbf{F}^{-1}(y) \right) - \psi(\tilde{\boldsymbol{\theta}}) + \ln \left(q^* \right) \right\} dP_0 \circ \mathbf{F}^{-1}(y)$ is the expectation parametrisa-

tion of the truncated exponential probability distribution derived from $P_{\tilde{\boldsymbol{\theta}}}$'s $(q^* - 1)$ th q-quantile of the fitness \mathbf{F} and $\boldsymbol{\mu} = \int \mathbf{x} \exp\{\tilde{\boldsymbol{\theta}}^{\mathsf{T}} \mathbf{t} (\mathbf{x}) - \psi(\tilde{\boldsymbol{\theta}})\} dP_0$ is the expectation parametrisation of $P_{\tilde{\boldsymbol{\theta}}}$.

Since varying q influences the magnitude of Eq. (3), as visualized in Fig. (2), we introduce an adaptive quantile into the optimization schedule by

$$q_{t+1} = q_t \exp\{\beta \cos(\alpha^t)\},\,$$

where α^t is the angle between three consecutive updates of Eq. 2 considered as points on the statistical manifold S. Parameter q_t can be regarded as an equivalent of the temperature parameter in the simulated annealing schedule.

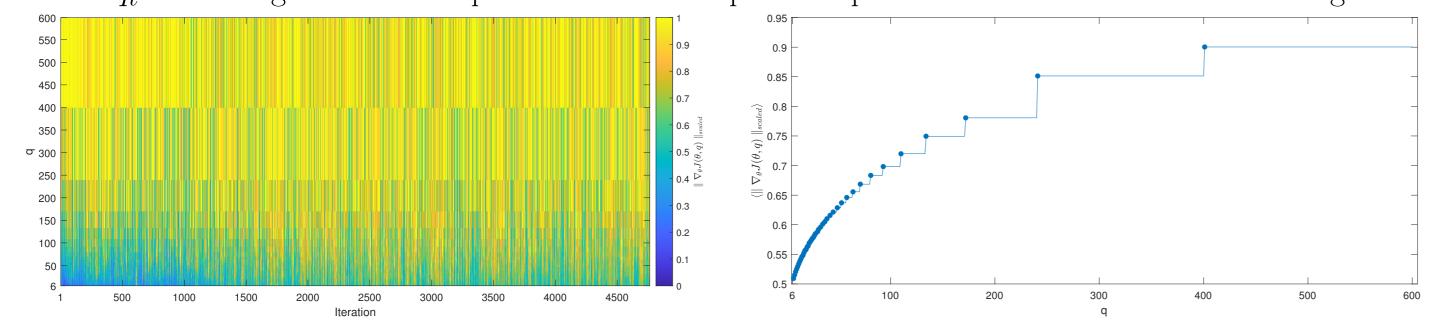


Figure 2: Influence of the q-quantile on the fitness gradient (Eq. (3)) in one run of the ETRPA on the problem of densest p2-packing of a regular octagon. (Left) Relationship between selection quantile q and the scaled expected fitness gradient $\|\nabla_{\boldsymbol{\theta}} J(\boldsymbol{\theta}^t, q)\|_{\text{scaled}} = \|\boldsymbol{\mu}_{\mathbf{F}}^t\|_{1}$

 $\frac{1-\bar{q}}{\|\boldsymbol{\mu}_{\mathbf{F}_{1-\frac{1}{q^{\max}}}}^{t}-\boldsymbol{\mu}^{t}\|_{1}}, \text{ where } q^{\max} = \operatorname{argmax}_{q} \parallel \boldsymbol{\mu}_{\mathbf{F}_{1-\frac{1}{q}}}^{t}-\boldsymbol{\mu}^{t} \parallel_{1}, \ q=6,\ldots,600. \ (\mathbf{Right}) \text{ Scaled expected fitness averaged through all iterations} \\ <\parallel \nabla_{\theta}J(\theta,q)\parallel_{scaled}> \text{ where markers denote changes in the value}.$









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Extended multivariate von Mises distribution

The lattice subgroup L of a CSG induces quotient space \mathbb{R}^n/L . Thus, we restrict ETRPA's search to a statistical model consisting of probability distributions defined on an n-torus. This is done by extending the multivariate von Mises model [4] to the family of distributions with the probability density function

$$f(\boldsymbol{\theta}|\boldsymbol{\mu}, \boldsymbol{\kappa}, \mathbf{D}) = \frac{1}{Z(\boldsymbol{\mu}, \boldsymbol{\kappa}, \mathbf{D})} \exp \left\{ \boldsymbol{\kappa}^{\mathsf{T}} c(\boldsymbol{\theta} - \boldsymbol{\mu}) + \frac{1}{2} \begin{bmatrix} c(\boldsymbol{\theta} - \boldsymbol{\mu}) \\ s(\boldsymbol{\theta} - \boldsymbol{\mu}) \end{bmatrix}^{\mathsf{T}} \mathbf{D} \begin{bmatrix} c(\boldsymbol{\theta} - \boldsymbol{\mu}) \\ s(\boldsymbol{\theta} - \boldsymbol{\mu}) \end{bmatrix} \right\}$$
(4)

$$c(\boldsymbol{\theta} - \boldsymbol{\mu}) = [cos(\theta_1 - \mu_1), \dots, cos(\theta_n - \mu_n)]^{\mathsf{T}}$$

$$s(\boldsymbol{\theta} - \boldsymbol{\mu}) = [sin(\theta_1 - \mu_1), \dots, sin(\theta_n - \mu_n)]^{\mathsf{T}}$$

$$0 \le \theta_i \le 2\pi \quad 0 \le \mu_i \le 2\pi \quad 0 \le \kappa_i,$$

$$\mathbf{D} = \begin{bmatrix} \mathbf{D}^{cc} & \mathbf{D}^{cs} \\ \mathbf{D}^{cs} & \mathbf{D}^{ss} \end{bmatrix}$$

where the diagonal elements of \mathbf{D}^{cc} , \mathbf{D}^{ss} and \mathbf{D}^{cs} are zero and \mathbf{D} is a $2n \times 2n$ real-valued symmetric matrix that controls cosine-cosine, sine-sine and cosine-sine interactions. Additionally, we remove non-identifiability of Eq. (4) by the restriction $\kappa_i > 0$ and introduce the exponential family re-parametrisation of the extended multivariate von Mises distribution Eq.(4) to

$$f(\boldsymbol{\theta}|\boldsymbol{\eta}, \mathbf{E}) = \exp\left\{ \begin{bmatrix} c(\boldsymbol{\theta}) \\ s(\boldsymbol{\theta}) \end{bmatrix}^{\mathsf{T}} \boldsymbol{\eta} + \operatorname{vec}\left(\begin{bmatrix} c(\boldsymbol{\theta}) \\ s(\boldsymbol{\theta}) \end{bmatrix} \begin{bmatrix} c(\boldsymbol{\theta}) \\ s(\boldsymbol{\theta}) \end{bmatrix}^{\mathsf{T}} \right)^{\mathsf{T}} \operatorname{vec}\left(\mathbf{E}\right) - \psi\left(\boldsymbol{\eta}, \mathbf{E}\right) \right\}$$

where $\text{vec}(\cdot)$ denotes vectorization.

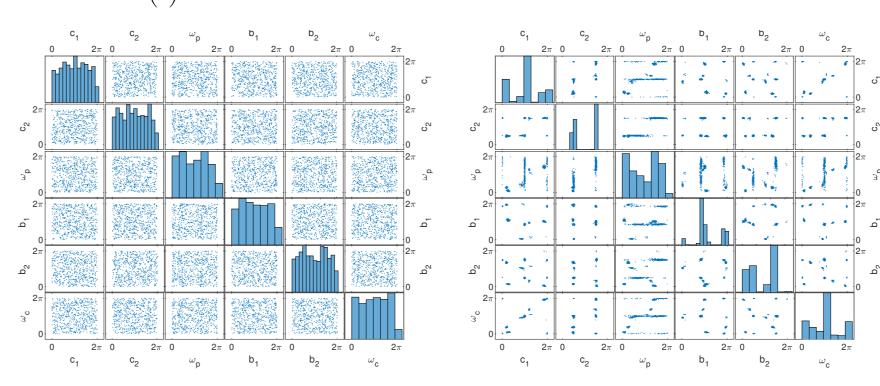


Figure 3: 2-dimensional projections along coordinate axes and histograms of univariate marginals corresponding to the respective optimization variables of 600 realizations of the exponential re-parametrisation of the extended multivariate von Mises distribution defined on an 6-torus. (Left) Initial distribution and (Right) output distribution of a single run of the ETRPA on the problem of densest p2-packing of a regular octagon with following optimization variables: octagon centroid fractional coordinates c_1 and c_2 in the p2 fundamental region, angle of rotation of the octagon ω_p , lengths of the lattice generators b_1 and b_2 , and angle between lattice generators ω_c .

Geometry of ETRPA

Recasting ETRPA as a solution of a maximin problem, which has an interpretation in terms of maximizing multiinformation [2] and can be observed visually in Fig. 3, allows us to view ETRPA (information) geometrically as follows. The aforementioned maximin problem reads

$$\max_{P_{\boldsymbol{\theta}^{q^*}} \in S^{q^*}} \min_{P_{\boldsymbol{\theta}^1} \in S^1} D_{KL}(P_{\boldsymbol{\theta}^{q^*}} \mid\mid P_{\boldsymbol{\theta}^1}); \ S^{q^*}, S^1 \subset S$$

$$(5)$$

and is induced by dual gradient flows between two codimension 1 submanifolds of the ambient statistical manifold

$$S = \bigcup_{q \in [1,\infty)} S^q, \ S^q = \{ dP^q(\boldsymbol{\theta}) | \boldsymbol{\theta} \in \boldsymbol{\Theta}^q \subset \mathbb{R}^n \},$$

where $dP^{q}(\boldsymbol{\theta}) = \begin{cases} \exp\left\{\boldsymbol{\theta}^{\intercal}\mathbf{t}\left(\mathbf{F}^{-1}(y)\right) - \psi\left(\boldsymbol{\theta}\right) + \ln(q)\right\} dP_{0} \circ \mathbf{F}^{-1}(y) & y \geq \mathbf{F}^{\boldsymbol{\theta}}_{1-\frac{1}{q}} \\ 0 & \text{otherwise} \end{cases}$ Indeed, the Euclidean fitness gradient Eq. (3) for a fixed q^{*} determines proximal map maximization [5]

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$$q^*$$
 determines proximal map maximizat
$$\boldsymbol{\theta}^{t+1} = \operatorname*{argmax}_{\boldsymbol{\theta} \in \boldsymbol{\Theta}} \left\{ \boldsymbol{\theta}^\intercal \nabla_{\boldsymbol{\theta}} J(\boldsymbol{\theta}^t) - \frac{1}{\epsilon} D_{KL} \left(P_{\boldsymbol{\theta}^t} \mid\mid P_{\boldsymbol{\theta}} \right) \right\},$$

and hence natural gradient updates in dual coordinate systems

$$\boldsymbol{\theta}^{t+1} = \boldsymbol{\theta}^t + \epsilon^t \mathcal{I}_{\boldsymbol{\theta}^t}^{-1} \nabla_{\boldsymbol{\theta}} J(\boldsymbol{\theta}^t), \ \boldsymbol{\mu}^{t+1} = \boldsymbol{\mu}^t + \epsilon^t \mathcal{I}_{\boldsymbol{\mu}^t}^{-1} \nabla_{\boldsymbol{\mu}} J(\boldsymbol{\mu}^t),$$
 (6)

which when lifted to S via $\boldsymbol{\theta}^1 = [\boldsymbol{\theta}, 1]$ and $\boldsymbol{\mu}^{q^*} = \left[\boldsymbol{\mu}_{\mathbf{F}_{1-\frac{1}{q^*}}}, -\frac{1}{q^*}\right]$, where $\boldsymbol{\mu}_{\mathbf{F}_{1-\frac{1}{q^*}}} = q^*\boldsymbol{\mu} - (q^*-1)\boldsymbol{\mu}_{\mathbf{F}_{\frac{1}{q^*}}}$, for the natural gradients of Eq. (6) in $\boldsymbol{\theta}$ and $\boldsymbol{\mu}$ coordinates respectively, give the flows. Moreover, for $\epsilon^t = \frac{\Delta^t}{\sqrt{\nabla_{\boldsymbol{\theta}} J(\boldsymbol{\theta}^t)^{\mathsf{T}}(\mathcal{I}_{\boldsymbol{\theta}^t})^{-1}\nabla_{\boldsymbol{\theta}} J(\boldsymbol{\theta}^t)}}$

the $\boldsymbol{\theta}$ -coordinate natural gradient update part of Eq. (6) coincides with trust region update Eq. (2). In addition, submanifolds S^q induce a dually flat structure on S and following the information projection theorems, solving Eq. (5) is equivalent to solving

$$\max_{P_{\boldsymbol{q}^*} \in S^{q^*}} D_{KL}(P_{\boldsymbol{\theta}^{q^*}} \parallel P_{\hat{\boldsymbol{\theta}}_s^1}) \tag{7}$$

where $P_{\hat{\boldsymbol{\theta}}_s^1} \in S^1$ is a model without any interaction whatsoever corresponding to the maximum entropy estimate of $P_{\boldsymbol{\theta}^{q^*}} \in S^{q^*}$. Consequently, $D_{KL}(\cdot \mid\mid P_{\hat{\boldsymbol{\theta}}_s^1})$ in Eq.(7) is considered as a special case of multi-information [6], a measure of stochastic dependence in complex systems.

Further information

M. TORDA, J. Y. GOULERMAS, R. Púček, AND V. Kurlin, Entropic Trust Region for Densest Crystallographic Symmetry Group Packings, SIAM Journal on Scientific Computing, 45.4 (2023), pp. B493-B522.

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